

University of Groningen

Large-scale structure in the HI Parkes All-Sky Survey

Basilakos, S.; Plionis, M.; Kovac, K.; Voglis, N.

Published in:
Monthly Notices of the Royal Astronomical Society

DOI:
[10.1111/j.1365-2966.2007.11781.x](https://doi.org/10.1111/j.1365-2966.2007.11781.x)

IMPORTANT NOTE: You are advised to consult the publisher's version (publisher's PDF) if you wish to cite from it. Please check the document version below.

Document Version
Publisher's PDF, also known as Version of record

Publication date:
2007

[Link to publication in University of Groningen/UMCG research database](#)

Citation for published version (APA):

Basilakos, S., Plionis, M., Kovac, K., & Voglis, N. (2007). Large-scale structure in the HI Parkes All-Sky Survey: filling the voids with HI galaxies? *Monthly Notices of the Royal Astronomical Society*, 378(1), 301-308. <https://doi.org/10.1111/j.1365-2966.2007.11781.x>

Copyright

Other than for strictly personal use, it is not permitted to download or to forward/distribute the text or part of it without the consent of the author(s) and/or copyright holder(s), unless the work is under an open content license (like Creative Commons).

The publication may also be distributed here under the terms of Article 25fa of the Dutch Copyright Act, indicated by the "Taverne" license. More information can be found on the University of Groningen website: <https://www.rug.nl/library/open-access/self-archiving-pure/taverne-amendment>.

Take-down policy

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

Downloaded from the University of Groningen/UMCG research database (Pure): <http://www.rug.nl/research/portal>. For technical reasons the number of authors shown on this cover page is limited to 10 maximum.

Large-scale structure in the H I Parkes All-Sky Survey: filling the voids with H I galaxies?

S. Basilakos,^{1*} M. Plionis,^{2,3} K. Kovač⁴ and N. Voglis^{1†}

¹*Academy of Athens, Research Centre for Astronomy & Applied Mathematics, Soranou Efessiou 4, 11-527 Athens, Greece*

²*Institute of Astronomy & Astrophysics, National Observatory of Athens, I. Metaxa & B. Pavlou, Palaia Penteli, 152 36 Athens, Greece*

³*Instituto Nacional de Astrofísica, Óptica y Electrónica (INAOE) Apartado Postal 51 y 216, 72000 Puebla, Pue., México*

⁴*Kapteyn Astronomical Institute, University of Groningen, the Netherlands*

Accepted 2007 March 23. Received 2007 March 21; in original form 2006 December 4

ABSTRACT

We estimate the two-point correlation function in redshift space of the recently compiled H I Parkes All-Sky Survey neutral hydrogen (H I) sources catalogue, which if modelled as a power law, $\xi(r) = (r_0/r)^\gamma$, the best-fitting parameters for the H I selected galaxies are found to be $r_0 = 3.3 \pm 0.3 h^{-1}$ Mpc with $\gamma = 1.38 \pm 0.24$. Fixing the slope to its universal value $\gamma = 1.8$, we obtain $r_0 = 3.2 \pm 0.2 h^{-1}$ Mpc.

Comparing the measured two-point correlation function with the predictions of the concordance cosmological model ($\Omega_\Lambda = 0.74$), we find that at the present epoch the H I selected galaxies are antibiased with respect to the underlying matter fluctuation field with their bias value being $b_0 \simeq 0.68$. Furthermore, dividing the H I galaxies into two richness subsamples we find that the low-mass H I galaxies have a very low present bias factor ($b_0 \simeq 0.48$), while the high-mass H I galaxies trace the underlying matter distribution as the optical galaxies ($b_0 \simeq 1$). Using our derived present-day H I galaxy bias we estimate their redshift-space distortion parameter, and correct accordingly the correlation function for peculiar motions. The resulting real-space correlation length is $r_0^{\text{re}} = 1.8 \pm 0.2 h^{-1}$ Mpc and $r_0^{\text{re}} = 3.9 \pm 0.6 h^{-1}$ Mpc for the low- and high-mass H I galaxies, respectively. The low-mass H I galaxies appear to have the lowest correlation length among all extragalactic populations studied to date. In order to corroborate these results we have correlated the IRAS–Point Source Catalogue for Redshift (PSCz) reconstructed density field, smoothed over scales of $5 h^{-1}$ Mpc, with the positions of the H I galaxies, to find that indeed the H I galaxies are typically found in negative overdensity regions ($\delta\rho/\rho_{\text{PSCz}} \lesssim 0$), even more so the low-mass H I galaxies.

Finally, we also study the redshift evolution of the H I galaxy linear bias factor and find that the H I-galaxy population is antibiased up to $z \sim 1.3$. While at large redshifts $z \sim 3$, we predict that the H I galaxies are strongly biased. Our bias evolution predictions are consistent with the observational bias results of Ly α galaxies.

Key words: galaxies: clusters: general – cosmology: theory – large-scale structure of Universe.

1 INTRODUCTION

The clustering of the different extragalactic sources is an ideal tool for testing theories of structure formation as well as studying the large-scale structure (e.g. Peebles 1993). The traditional indicator of clustering, the two-point correlation function, is a fundamental statistical test for the study of the extragalactic distribution and is relatively straightforward to measure from observational data. Re-

cently, the new generation redshift surveys such as the Point Source Catalogue for Redshift (PSC-z; Saunders et al. 2000), the European Large-Area ISO survey (ELAIS; Oliver et al. 2000), the Sloan Digital Sky Survey (SDSS; York 2000) the 2dF Galaxy Redshift Survey (2dFGRS; Colless et al. 2001), the DEEP2 Galaxy Redshift Survey (DEEP2; Davis et al. 2003) and the VIMOS-VLT Deep Survey (VVDS; Le Fèvre et al. 2005) have extended the clustering studies and now we are able to extract definitive measurements of galaxy clustering in the local and relatively distant universe.

Indeed Zehavi et al. (2002) and Hawkins et al. (2003) have found $r_0 \simeq 5 h^{-1}$ Mpc and $\gamma \simeq 1.7$ for the SDSS galaxies, while Norberg et al. (2001) found for the 2dFGRS galaxies that $r_0 \simeq 4.9 h^{-1}$ Mpc

*E-mail: svasil@academyofathens.gr

†Deceased, 2007 February 9.

and $\gamma \simeq 1.7$. Also, Coil et al. (2006) have found $r_0 \simeq 4 h^{-1}$ Mpc and $\gamma \simeq 1.7$ for the DEEP2 galaxies at $z \sim 1$ and Guzzo et al. (2007) derived $r_0 \simeq 3.6 h^{-1}$ Mpc for the VVDS galaxies at $z \sim 1.5$. In the infrared side, Jing, Börner & Suto (2002) derived $r_0 \simeq 3.7 h^{-1}$ Mpc and $\gamma \simeq 1.7$, using the IRAS–PSCz data, while Gonzalez-Solares et al. (2004) using the ELAIS sources found $r_0 \simeq 4.3 h^{-1}$ Mpc and $\gamma \simeq 2$ correlation functions. From a cosmological point of view, it is imperative to understand how the different types of galaxies trace the underlying mass distribution. It is well known that the large-scale clustering pattern of different mass tracers (galaxies, active galactic nuclei, clusters, etc.) is characterized by a bias picture (Kaiser 1984; Bardeen et al. 1986). In particular, biasing is assumed to be statistical in nature; galaxies and clusters are identified as high peaks of an underlying, initially Gaussian, random density field. Biasing of galaxies with respect to the dark matter distribution was also found to be an essential ingredient of cold dark matter (CDM) models of galaxy formation in order to reproduce the observed galaxy distribution (e.g. Benson et al. 2000).

In this paper we utilize the recently completed H I Parkes All-Sky Survey (HIPASS; Barnes et al. 2001), which is the largest uniform sample of H I selected galaxies in the local universe, attempting to make a detailed investigation of the connection between the clustering and the biasing properties of H I selected galaxies. In particular we determine, in the framework of the concordance Λ CDM cosmology, the relative H I galaxy bias at the present time and describe the corresponding bias as a function of redshift. For a detailed study of the observational clustering properties of the HIPASS sources we refer the reader to the recent works of Ryan-Weber (2006) and Meyer et al. (2007). We also refer the reader to the new *Arecibo Legacy Fast ALFA* (ALFALFA) drift survey, especially designed to probe the faint end of the H I mass function in the local universe (Giovanelli et al. 2005a,b).

The structure of the paper is as follows. In Section 2 we discuss the H I galaxy data set, its measured redshift-space correlation function, the determination of its present-day bias factor, the corrected for redshift-space distortion real-space correlation function as well as an investigation of the environment of H I galaxies. In Section 3 we present two bias evolution models and study the corresponding evolution for the case of H I galaxies within the concordance Λ CDM model. Finally, we draw our conclusions in Section 4.

2 ESTIMATION OF THE H I SELECTED GALAXIES CORRELATION FUNCTION

2.1 The HIPASS data

The recent HIPASS catalogue (for more details see Meyer et al. 2004; Ryan-Weber 2006), contains 4315 neutral hydrogen (H I) sources in the southern region of the survey ($\delta < +2^\circ$), covering an area of $\sim 2.07\pi$ on the sky between 300 km s^{-1} and $12\,700 \text{ km s}^{-1}$ (see also Doyle et al. 2005). We exclude detections of which the expected completeness is very low ($C < 0.5$) in order to avoid introducing systematic effects related to the H I galaxy space density (estimation of the completeness limit C is given in Zwaan et al. 2004, 2005a). The remaining sample contains ~ 4008 H I selected galaxies with masses $M_{\text{H I}} \leq 4 \times 10^{10} h^{-2} M_\odot$.¹ In the meantime northern extension of the HIPASS has been published

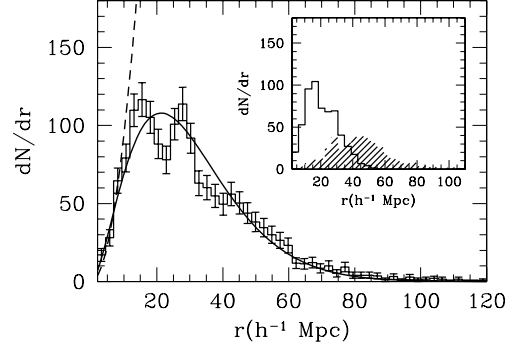


Figure 1. The measured (histogram) and the fitted (continuous line) number of the H I galaxy sources as a function of distance. The dashed line corresponds to a volume-limited expectation. The insert panel shows the corresponding selection function for the S_1 (solid) and the S_2 (hatched) sub-samples, respectively (see Section 2.2).

containing 1002 H I detections in the following area $2^\circ < \delta < 20^\circ$ (Wong et al. 2006).

The measured radial velocities are translated to the Local group frame and are converted to proper distances using a spatially flat cosmology with $H_0 = 100 h \text{ km s}^{-1} \text{ Mpc}^{-1}$ and $\Omega_m = 1 - \Omega_\Lambda = 0.26$.

In Fig. 1, we present the estimated (histogram) and that expected for a volume-limited sample (dashed line), number of the H I sources as a function of distance. It appears that the H I galaxy distribution is roughly volume-limited out to $10 h^{-1}$ Mpc, a fact corroborated also by the χ^2 minimization test between model and observations, which gives a reduced $\chi^2/\text{d.o.f.}$ (up to $r \leq 10 h^{-1}$ Mpc) of 1.04. However, due to the fact that the HIPASS catalogue is a peak flux-limited sample it suffers from the well known degradation of sampling as a function of distance (codified by the so-called *selection function*). Thus, we can attempt to parametrize the H I galaxy distance distribution, for this particular effect, using the following formula:

$$\frac{dN}{dr} = d\Omega r^2 \langle \rho \rangle \phi(r), \quad (1)$$

where $\phi(r)$ is the parametrized selection function

$$\phi(r) = \frac{A}{d\Omega \langle \rho \rangle} r^\alpha \exp \left[- \left(\frac{r}{r_s} \right)^m \right] \quad (2)$$

with

$$A = \frac{Nm}{\Gamma[(\alpha + 3)/m] r_s^{\alpha+3}}. \quad (3)$$

Note that N is the total number of the H I sources, $d\Omega$ is the solid angle of the survey, $\langle \rho \rangle$ is the mean density of the HIPASS sample (see below) and the factor Γ is the Gamma function. The solid line in Fig. 1 corresponds to the best-fitting dN/dr , which is determined by the standard χ^2 minimization procedure in which each part of the histogram is weighted by its Poisson error. The reduced $\chi^2/\text{d.o.f.}$ is ~ 0.90 and the corresponding values of the fitted parameters are $r_s = 16.5^{+1.0}_{-2.5} h^{-1} \text{ Mpc}$, $m = 1.24^{+0.06}_{-0.08}$ and $\alpha = -0.30 \pm 0.12$ (the uncertainties correspond to 1σ errors). In Fig. 2 we present the number density of our sample as a function of distance. Up to $\sim 10 h^{-1}$ Mpc the corresponding space density is roughly constant, being $\langle \rho \rangle = 0.14 \pm 0.02 h^3 \text{ Mpc}^{-3}$ and then drops because of the sample flux limit.

¹ Zwaan et al. (2005a) used $H_0 = 75 h_{75} \text{ km s}^{-1} \text{ Mpc}^{-1}$. In that case the maximum H I mass is $\sim 7.1 \times 10^{10} h_{75}^{-2} M_\odot$.

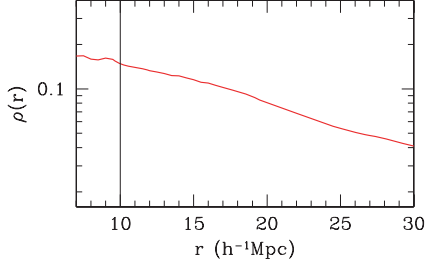


Figure 2. The space density of the HIPASS galaxy data as a function of distance.

2.2 H I selected galaxy redshift-space correlation function

We estimate the redshift-space H I galaxy correlation function using the estimator described by Efstathiou et al. (1991)

$$\xi_z(r) = f \frac{N_{DD}}{N_{DR}} - 1, \quad (4)$$

where N_{DD} is the number of H I pairs in the interval $[r - \Delta r, r + \Delta r]$, and N_{DR} is the number of data-random pairs. In the above relation the normalization factor is $f = 2N_R/(N_D - 1)$, with N_D and N_R the total number of data and random points, respectively. The random distributions were constructed using 100 Monte Carlo realizations, respecting the survey area and taking into account all possible systematic effects of the data (e.g. completeness, fraction of H I sources missed by the finding algorithm due to the HIPASS flux limit and the redshift distribution of the H I galaxy data). Note, that we have tested that our random catalogues reproduce exactly the corresponding H I selection function.

In the left-hand panel of Fig. 3, we present the measured redshift-space two-point correlation function in logarithmic intervals (dots). The clustering is evident, although the correlation function beyond $r \geq 11 h^{-1}$ Mpc drops dramatically. The dashed line corresponds to the best-fitting power-law model $\xi(r) = (r_0/r)^\gamma$, determined by the standard χ^2 minimization procedure, in which each correlation point is weighted by the inverse of its error.

In Fig. 4 we present the iso- $\Delta\chi^2$ contours (where $\Delta\chi^2 = \chi^2 - \chi^2_{\min}$) in the $r_0 - \gamma$ plane (see the index All). The χ^2_{\min} is the absolute minimum value of the χ^2 . In Table 1 we present the best-fitting parameters, derived for $r \geq 1 h^{-1}$ Mpc in order to avoid the signal from the smallest and therefore the highly non-linear scales (e.g. Vasiliev,

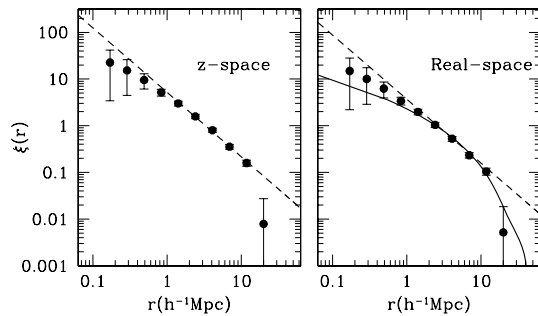


Figure 3. Left-hand panel: the HIPASS galaxy spatial two-point correlation function (points) in redshift space. The uncertainties were estimated using 100 bootstrap resamplings of the data (Mo, Jing & Börner 1992). Right-hand panel: the HIPASS correlation function in real space. The dashed line represents the best-fitting power law $\xi(r) = (r_0/r)^\gamma$ (see parameters in Table 1), while the continuous line represents the best-fitting Λ CDM ($\Omega_\Lambda = 0.74$) model with $b_0 = 0.68$.

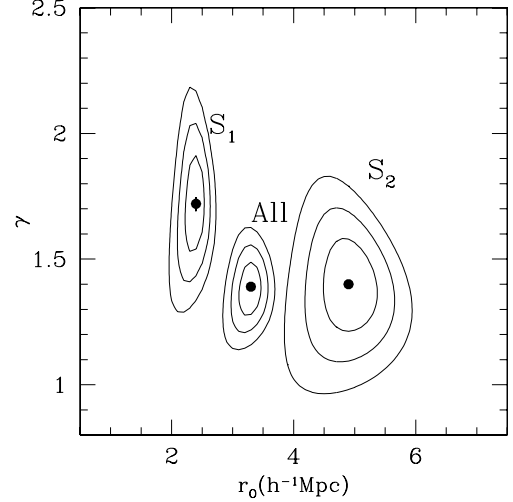


Figure 4. Iso- $\Delta\chi^2$ contours in the $r_0 - \gamma$ parameter space for the whole catalogue and separately for the S_1 and S_2 samples. The contours correspond to 1σ ($\Delta\chi^2 = 2.30$), 2σ ($\Delta\chi^2 = 6.17$) and 3σ ($\Delta\chi^2 = 11.8$) uncertainties, respectively.

Baryshev & Sylos Labini 2006). Note that our results remain robust by varying the upper r limit within the $5\text{--}20 h^{-1}$ Mpc range.

The resulting clustering parameters² are $r_0 = 3.3 \pm 0.3 h^{-1}$ Mpc and $\gamma = 1.38 \pm 0.24$. Fixing now the correlation function slope to its nominal value of $\gamma = 1.8$, we obtain $r_0 = 3.2 \pm 0.2 h^{-1}$ Mpc, but now the fit is good for $r \geq 2 h^{-1}$ Mpc. Note that our results are in good agreement with a similar analysis of the HIPASS data by Ryan-Weber (2006), who find $r_0 = 3.1 \pm 0.5 h^{-1}$ Mpc with $\gamma = 1.4 \pm 0.5$.

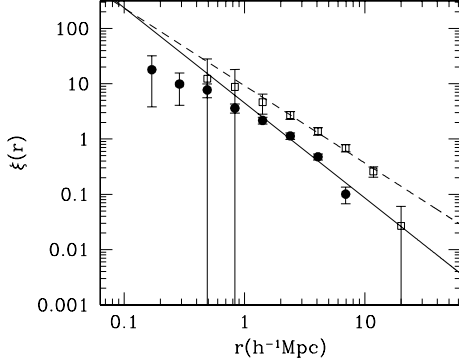
It is interesting to compare our results with other recent galaxy correlation function determinations in the local universe. In the optical, Zehavi et al. (2002) and Hawkins et al. (2003) find $r_0 \simeq 5 h^{-1}$ Mpc and $\gamma \simeq 1.7$ for the SDSS galaxies, while Norberg et al. (2001) find $r_0 \simeq 4.9 h^{-1}$ Mpc and $\gamma \simeq 1.7$ for the 2dFGRS galaxies. In the infrared, Jing et al. (2002) find for the IRAS-PSCz data $r_0 \simeq 3.7 h^{-1}$ Mpc and $\gamma \simeq 1.7$, while Gonzalez-Solares et al. (2004) using the ELAIS sources found $r_0 \simeq 4.3 h^{-1}$ Mpc and $\gamma \simeq 2$. It should be pointed out that the HIPASS galaxy correlation function has a significantly shallower slope than the optical and infrared galaxy correlation functions.

In order to investigate the possible H I mass dependence of the correlation function, we divided the sample into two subsamples using the following mass thresholds: (i) $M_{H I} \leq 2.475 \times 10^9 h^{-2} M_\odot$, hereafter S_1 sample containing 2389 entries, and (ii) $M_{H I} > 2.475 \times 10^9 h^{-2} M_\odot$, hereafter S_2 sample containing 1619 entries. We apply our statistical analysis to each H I subsample by taking into account the individual redshift selection functions (see insert of Fig. 1). The resulting correlation function parameters are presented in the last two rows of Table 1 and in Fig. 4, where we show the iso- $\Delta\chi^2$ contours for both the S_1 and S_2 cases, separately. In Fig. 5, we present the estimated two-point redshift correlation function for the S_1 (solid circles) and S_2 (open squares) subsamples, respectively. The lines correspond to the best-fitting power-law model. In particular, for the low-mass H I galaxies (S_1 sample) we obtain $r_0 = 2.4 \pm 0.3 h^{-1}$ Mpc

² The robustness of our results to the fitting procedure was tested using different numbers of bins (spanning from 10 to 20) and we found no significant differences.

Table 1. Results of the correlation function analysis of the HIPASS sample. Errors of the fitted parameters represent 3σ uncertainties. Finally, the r_0 has units of h^{-1} Mpc.

Sample	N	r_0	γ	$r_0(\gamma = 1.8)$	b_0	β	$K(\beta)$	r_0^{re}	$r_0^{\text{re}}(\gamma = 1.8)$
All	4008	3.3 ± 0.3	1.38 ± 0.24	3.2 ± 0.2	0.68 ± 0.10	0.66	1.52	2.4 ± 0.3	2.5 ± 0.2
S_1	2389	2.4 ± 0.3	1.72 ± 0.46	2.4 ± 0.2	0.48 ± 0.10	0.93	1.79	1.8 ± 0.3	1.8 ± 0.2
S_2	1619	4.9 ± 1.0	1.40 ± 0.42	4.6 ± 0.7	0.94 ± 0.15	0.47	1.36	4.0 ± 0.9	3.9 ± 0.6

**Figure 5.** The spatial two-point correlation function in redshift space for the S_1 (solid points) and S_2 (open squares) subsamples. The error bars are estimated using the bootstrap procedure. The lines represent the best-fitting power-law model, $\xi(r) = (r_0/r)^\gamma$ (see parameters in Table 1).

and $\gamma = 1.72 \pm 0.46$ (continuous line). While, for the high-mass H I galaxies (S_2 case) we find a larger correlation length, $r_0 = 4.9 \pm 1.0 h^{-1}$ Mpc and $\gamma = 1.40 \pm 0.42$ (see dashed line in Fig. 5). It is clear that the correlation length increases with richness, as expected from its well known richness dependence of the correlation function (e.g. Bahcall & Burgett 1986). Note that in Section 2.4 we further investigate this issue in more detail.

2.3 The H I selected galaxy real-space correlation function

The previous results are hampered by the fact that the analysis has been performed in redshift space and therefore the derived correlation functions are amplified in the linear regime by the factor $K(\beta)$ (Hamilton 1992, also see Hawkins et al. 2003) given by

$$K(\beta) = 1 + \frac{2\beta}{3} + \frac{\beta^2}{5}, \quad (5)$$

where $\beta \simeq \Omega_m^{0.6}/b_0$, and b_0 is the present-day bias factor of the H I selected galaxies. Determining this bias factor and using $\Omega_m = 0.26$, valid for the concordance cosmological model, we can evaluate the $K(\beta)$ factor and recover the real-space H I galaxy correlation function.

To this end, we compare the theoretically expected mass-tracer correlation function, in the concordance CDM cosmology, with our derived correlation function. It is well known (e.g. Kaiser 1984; Benson et al. 2000) that according to the linear biasing ansatz, the correlation functions of a mass-tracer (ξ_{obj}) and the dark matter (ξ_{DM}), are related by

$$\xi_{\text{obj}}(r) = b_0^2 \xi_{\text{DM}}(r). \quad (6)$$

We quantify the clustering of the underlying mass distribution using the spatial correlation function of the mass, $\xi_{\text{DM}}(r)$, which is the

Fourier transform of the spatial power spectrum $P(k)$:

$$\xi_{\text{DM}}(r) = \frac{1}{2\pi^2} \int_0^\infty k^2 P(k) \frac{\sin(kr)}{kr} dk, \quad (7)$$

where k is the comoving wavenumber in units of h Mpc $^{-1}$. As for the power spectrum, we consider that of CDM models: $P(k) \approx k^n T^2(k)$, with scale-invariant ($n = 1$) primeval inflationary fluctuations. We utilize the transfer function parameterization as in Bardeen et al. (1986), with the approximate corrections given by Sugiyama's (1995) formula:

$$T(k) = \frac{\ln(1 + 2.34q)}{2.34q} [1 + 3.89q + (16.1q)^2 + (5.46q)^3 + (6.71q)^4]^{-1/4}$$

with

$$q = \frac{k}{\Omega_m h^2 \exp[-\Omega_b - (2h)^{1/2} \Omega_b / \Omega_m]}, \quad (8)$$

where Ω_b is the baryon density. Note that we also use the non-linear corrections introduced by Peacock & Dodds (1994).

In the present analysis we consider the concordance model ($\Omega_m = 0.26$) with cosmological parameters, that fit the majority of observations, $\Omega_m + \Omega_\Lambda = 1$, $H_0 = 100 h$ km s $^{-1}$ Mpc $^{-1}$ with $h \simeq 0.72$ (cf. Freedman et al. 2001; Peebles & Ratra 2003; Spergel et al. 2003; Tonry et al. 2003; Riess et al. 2004; Tegmark et al. 2004; Basilakos & Plionis 2005, 2006a, and references therein), baryonic density parameter $\Omega_b h^2 \simeq 0.02$ (e.g. Olive, Steigman & Walker 2000; Kirkman et al. 2003). Note that the concordance model is normalized to have fluctuation amplitude, in spheres of $8 h^{-1}$ Mpc radius, of $\sigma_8 \simeq 0.74$ in agreement with both the recent *Wilkinson Microwave Anisotropy Probe* 3-yr results (Spergel et al. 2006) and the clustering of the XMM X-ray sources at relatively high redshifts (Basilakos & Plionis 2006a).

In order to derive the HIPASS bias at the present time we perform a standard χ^2 minimization procedure between the HIPASS galaxy correlation function in real space, $\xi_z(r)/K(\beta)$, and that expected in the concordance cosmological model:

$$\chi^2(b_0) = \sum_{i=1}^n \left[\frac{\xi_z^i(r)/K(\beta) - \xi_{\text{obj}}^i(r, b_0)}{\sigma^i} \right]^2, \quad (9)$$

where σ^i is the observed correlation function uncertainties. The outcome of this analysis provides the H I galaxy bias factor at the present time: $b_0 = 0.68 \pm 0.10$. Thus evidently the H I selected galaxies, in the local universe, are antibiased with respect to the underlying matter distribution, implying that they generally trace low-density regions. In the right-hand panel of Fig. 3, we plot the $\xi(r)$ in real space of the HIPASS sample and the theoretically expected one (continuous line) for the concordance model using $b_0 = 0.68$. It is obvious that they compare extremely well.

Performing now the same analysis for the two subsamples, we find for the S_1 subsample (low H I masses) a value of $b_0 \simeq 0.48$,

which again means that from the statistical point of view the low-mass H I galaxies trace low-density regions in the local universe. However, for the S_2 case we get $b_0 \simeq 0.94$, in agreement with that found using optical galaxies (e.g. Lahav et al. 2002; Verde 2002).

In Table 1 we list our derived values of the H I bias factor, b_0 , at the present time, as well as the redshift distortion β parameter and a measure of the $K(\beta)$ correction. Multiplying with $1/K(\beta)$ each bin of our redshift-space correlation function $\xi_z(r)$ and repeating our fitting procedure we derive an estimate of the real-space correlation length which is $r_0^e = 2.4 \pm 0.3 h^{-1}$ Mpc for $\gamma = 1.38$ and $r_0^e = 2.5 \pm 0.2 h^{-1}$ Mpc for $\gamma = 1.8$.

It is interesting to mention that Ryan-Weber (2006) and Meyer et al. (2007) found a somewhat larger correlation length in real space, $r_0^e = 3.45 \pm 0.25 h^{-1}$ Mpc with $\gamma = 1.47 \pm 0.08$ and $r_0^e = 3.5 \pm 0.7 h^{-1}$ Mpc with $\gamma = 1.9 \pm 0.3$, respectively. This difference should probably be attributed to the intrinsic uncertainty of the different methods used to correct the measured galaxy two-point correlation function for redshift-space distortions.

Note that for the S_1 and S_2 subsamples we find $r_0^e = 1.8 \pm 0.2 h^{-1}$ Mpc and $r_0^e = 3.9 \pm 0.6 h^{-1}$ Mpc for $\gamma = 1.8$. Finally, we would like to stress that the low-mass H I galaxies appear to have the lowest correlation length among all extragalactic populations studied to date.

2.4 The environment of H I galaxies

In order to investigate the cause for the relatively low-amplitude of the H I galaxy correlation function, we decided to correlate the positions of the H I galaxies with the reconstructed smoothed density field of the PSCz flux-limited (0.6 Jy) IRAS galaxy catalogue. We use the reconstructed density field of Branchini et al. (1999), smoothed as in Plionis & Basilakos (2001) utilizing a Gaussian kernel on a 3D grid with grid cell size and a smoothing scale (R_{sm}) equal to $5 h^{-1}$ Mpc. We also correct for effects related to the unavoidable convolution of the constant smoothing window with the degradation of the redshift selection function (see Plionis & Basilakos 2001, for details). We then cross-correlate the positions of each H I galaxy with the PSCz smoothed density field and assign to each the overdensity of its nearest grid cell. In this way we create the one-point PSCz overdensity probability distribution which is related to the H I galaxy positions and compare it with the corresponding full PSCz one-point distribution function. Note that we exclude from our analysis the inner volume ($r < 15 h^{-1}$ Mpc) in order to avoid discreteness effects, that is, associating a large number of HIPASS galaxies with only a few PSCz density grid cells (16 per cent of H I galaxies with 0.33 per cent of the total grid cells).

In the left-hand panel of Fig. 6 we present the two frequency distributions, for which the Kolmogorov–Smirnov two-sided test shows that the probability of them having the same parent distribution is $< 10^{-10}$. In the right-hand panel of the same figure we show the ratio of the two distributions as a function of $\delta\rho/\rho$. If the H I galaxies would trace the same structures as the IRAS–PSCz galaxies then the two distribution should be indistinguishable and the ratio would be ~ 1 . The results presented in this panel are very informative regarding the local environment of H I galaxies. We see that the H I galaxies avoid high-density regions, as well as the very low-density regions. They prefer, with respect to IRAS galaxies, regions of quite low galaxy density ($-0.4 \lesssim \delta\rho/\rho_{\text{PSCz}} \lesssim 0$).

In order to verify the other outcome of our H I galaxy correlation function analysis, that is, the fact that there appears to be a dichotomy between the clustering properties of the low- and high-mass subsamples of H I galaxies, we decided to compare the PSCz density

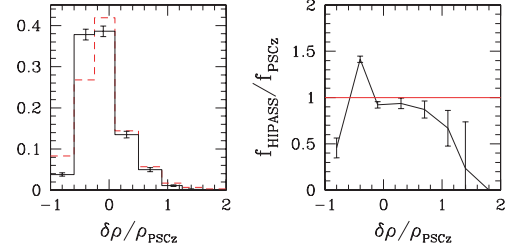


Figure 6. Left-hand panel: the one-point PSCz overdensity distribution function related to the positions of the H I galaxies (solid line) and the corresponding full PSCz distribution (dashed line). Right-hand panel: the ratio of the two distributions.

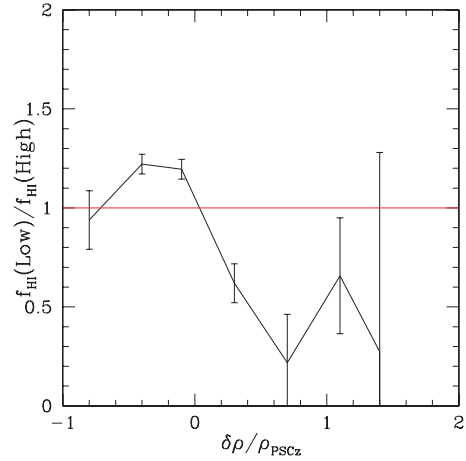


Figure 7. The ratio of the one-point PSCz overdensity distribution functions of the low- and high-mass H I galaxies.

distribution corresponding to the positions of the high- and low-mass H I galaxies. In Fig. 7 we show this ratio and it is evident that with respect to the high-mass H I galaxies, the low-mass ones populate typically only underdense void-like regions ($\delta\rho/\rho_{\text{PSCz}} \lesssim 0$). We find that 75 per cent of low-mass H I galaxies are associated with negative PSCz overdensities, with 40 per cent having $\delta\rho/\rho_{\text{PSCz}} < -0.3$. This fact corroborates our previous clustering analysis and the resulting extremely low-bias parameter of these H I galaxies ($b_0 \simeq 0.48$). It should also be kept in mind that with respect to the optical, infrared galaxies typically avoid the high-density regions, like those related to clusters of galaxies, a fact which translates to a PSCz correlation length of $r_0 \simeq 3.7 h^{-1}$ Mpc (Jing et al. 2002), while the PSCz bias parameter, within the concordance cosmological model ($\Omega_m = 0.26$), has been found to be $b_{\text{PSCz}} \simeq 0.9$ (e.g. Basilakos & Plionis 2006a).

Furthermore, these results are very interesting in themselves because they indicate that low-density void-like regions could well be filled by such H I low-mass galaxies, although it also appears that the lowest density regions are avoided also by such galaxies. We believe that this issue will be settled in the next few years with the new ALFALFA survey, which is expected to observe about 16 000 H I galaxies, most of them of low mass, in 7000 deg^2 (Giovannelli et al. 2005a,b).

3 H I GALAXIES BIAS EVOLUTION

In order to understand better the importance of the measured H I galaxy clustering, we investigate the evolution of the H I bias factor

as a function of redshift. The evolution of neutral hydrogen as a function of redshift is a powerful cosmological tool because it indicates the rate of transformation of gas into stars and thus the gas consumption and the evolution of star formation in the Universe.

Over the past two decades, based on different assumptions, a number of bias evolution models have been proposed (e.g. Nusser & Davis 1994; Fry 1996; Mo & White 1996; Matarrese et al. 1997; Tegmark & Peebles 1998; Basilakos & Plionis 2001). Here, we will use only models that have been shown to describe relatively well the evolution of bias even beyond $z \sim 1$. These are as follows.

(i) *Merging bias model* (hereafter B1): Mo & White (1996) have developed a model for the evolution of the so-called correlation bias, using the Press–Schechter formalism. Utilizing a similar formalism, Matarrese et al. (1997) extended the Mo & White (1996) results to include the effects of different mass scales (see also Bagla 1998). In this case the expression which describes the bias evolution is

$$b_{B1}(z) = 0.41 + \frac{(b_0 - 0.41)}{D^\beta(z)}, \quad (10)$$

with $\beta \simeq 1.8$. Note that $D(z)$ is the linear growth rate of clustering (cf. Peebles 1993)³ scaled to unity at the present time.

$$D(z) = \frac{5\Omega_m E(z)}{2} \int_z^\infty \frac{(1+x)}{E^3(x)} dx \quad (11)$$

with

$$E(z) = [\Omega_m(1+z)^3 + \Omega_\Lambda]^{1/2}. \quad (12)$$

(ii) *Bias from linear perturbation theory* (hereafter B2): Basilakos & Plionis (2001, 2003), using linear perturbation theory and the Friedmann–Lemaître solutions of the cosmological field equations have derived analytically the functional form for the evolution of the linear bias factor, b , between the background matter and a mass-tracer fluctuation field. The main assumptions of this model is that the DM and galaxies share the same velocity field and that the mass tracer population is conserved in time. For the case of a spatially flat Λ cosmological model ($\Omega_m + \Omega_\Lambda = 1$), the bias evolution can be written as

$$b_{B2}(z) = C_1 E(z) + C_2 E(z) I(z) + 1 \quad (13)$$

with

$$I(z) = \int_{1+z}^\infty \frac{y^3 dy}{(\Omega_m y^3 + \Omega_\Lambda)^{3/2}}. \quad (14)$$

Note that this model gives a family of bias curves, due to the fact that it has two unknown parameters, (the integration constants C_1, C_2). Basilakos & Plionis (2001, 2003) compared the B2 bias evolution model with other models as well as with the *Hubble Deep Field* biasing results (Arnouts et al. 2002), and found a good consistency. Of course in order to obtain partial solutions for $b(z)$ we need to estimate the values of the constants C_1 and C_2 , which means that we need to calibrate the $b(z)$ relation using two different epochs: $b(0) = b_0$ and $b(z_1) = b_1$.

Therefore, utilized the general bias solution (see equation 13), it is routine to obtain the expressions for the above constants as a function of b_0 and b_1 :

$$C_1 = \frac{(b_0 - 1)E(z_1)I(z_1) - (b_1 - 1)E(0)I(0)}{E(0)E(z_1)[I(z_1) - I(0)]}, \quad (15)$$

$$C_2 = \frac{E(0)(b_1 - 1) - E(z_1)(b_0 - 1)}{E(0)E(z_1)[I(z_1) - I(0)]}, \quad (16)$$

³ $D(z) = (1+z)^{-1}$ for an Einstein–de Sitter universe.

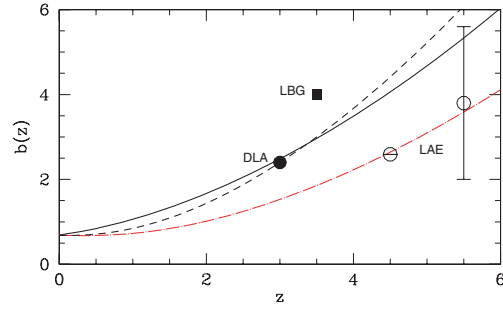


Figure 8. The evolution of the H I galaxy bias factor for the B1 (continuous line) and B2 (dashed line) bias models, respectively. The solid circle represents the observational bias for the DLA systems of Cooke et al. (2006). The square-like symbol corresponds to LBGs (Steidel et al. 1998) while the open circle represents the bias for the LAEs derived by Ouchi et al. (2005) and Kovač et al. (in preparation). Note that the dot–dashed line represents the bias behaviour for the LAE galaxies.

where for the present epoch we have: $b_0 \simeq 0.68$, $E(0) = 1$ and $I(0) \simeq 11.54$.

3.1 The H I bias factor in the distant universe

The hyperfine transition responsible for the H I emission is very weak and therefore with current telescopes, studies of the emission at 21 cm are limited to the very local universe ($z \leq 0.2$). This implies that the estimation of the bias parameter b_1 of H I sources in the distant universe ($z > 1.5$) would not be an easy task. At such redshifts, our knowledge on the neutral gas comes mostly from the absorption lines detected in the spectra of background quasars.

At redshifts $1.5 \leq z \leq 5.0$ most of the cosmic H I mass is contained in damped Ly α (DLA) systems (e.g. Wolfe et al. 1986; Lanzetta, Wolfe & Turnshek 1995). The DLA systems are quasar absorption line systems with H I column density $N_{\text{H I}} \geq 2 \times 10^{20}$ atoms cm⁻² (see Peroux et al. 2001; Wolfe, Gawiser & Prochaska 2005) and they are composed of predominantly neutral gas. Wolfe et al. (1986) established the given column density threshold so that it corresponds to the surface density limit of the local 21-cm observations at that time, which roughly corresponded to the transition from primarily ionized gas to predominantly neutral gas (e.g. Viegas 1995; Prochaska 1999; Vladilo et al. 2001). Therefore, these systems play a vital role in the structure formation because the large column densities can protect large amounts of neutral gas from the ionizing background (Zwaan & Prochaska 2006) and thus produce a favourable environment for star formation. To this end it is interesting to mention that Zwaan et al. (2005b) have found that in the redshift range of $z \sim 5$ to $z = 0$ the vast majority of the mass density in H I sources is locked up in column densities of $N_{\text{H I}} > 2 \times 10^{20}$ atom cm².

We will be using the above ideas to calibrate the B2 model, based on the Cooke et al. (2006) value of the bias for DLA systems which is $b(3) \simeq 2.4$ (see solid point in Fig. 8).⁴

3.2 The global H I bias results

Knowing the H I galaxy bias factor at the present time, we can use equations (15) and (16) and the observed DLA bias factor at $z \simeq 3$ to plot the B2 bias evolution model for the concordance cosmological

⁴ For $z = 3$ and $\Omega_m = 1 - \Omega_\Lambda = 0.26$ we get $E(3) \simeq 4.17$ and $I(3) \simeq 7.47$.

model (see Fig. 8). We also plot the B1 evolution model, which is uniquely determined from the present-day H I galaxy bias factor (B1 continuous line and B2 dashed line). Evidently, up to $z \leq 1.3$ we have an antibias picture but due to the fact that the bias is a monotonically increasing function of redshift, for both B1 and B2 biasing models, the H I sources at higher redshifts are biased with respect to the underlying matter distribution. Note also that both models provide almost the same bias evolution behaviour, while for $z \gtrsim 5$ the two models diverge.

We further compare our analytic solutions with observations of: (i) Lyman-break galaxies (hereafter LBGs; see solid square in Fig. 8) which are strongly biased with respect to the underlying mass and have $b(3.4) \simeq 4$ (Adelberger et al. 1998; Steidel et al. 1998; Kashikawa et al. 2006), and (ii) Ly α emitters (hereafter LAEs; see open circles in Fig. 8) with $b(5.7) = 3.8 \pm 1.8$ (Ouchi et al. 2005) and $b(4.5) \simeq 2.6$ (Kovač et al., in preparation). For the LAE case we present in Fig. 8 the B2 bias evolution model, anchored at $z = 4.5$ (dot-dashed line).⁵ Within this bias evolution model the LAE galaxies appear antibiased up to $z \simeq 2.1$, while beyond a redshift of $z \sim 3$ they become strongly biased with respect to the underlying mass fluctuations.

The strong biasing predictions, in agreement with those found from simulations of galaxy formation (e.g. Kauffmann et al. 1999), imply that Ly α systems, at high redshifts, are formed at the highest peaks of matter density field (e.g. Mo & White 2002). The last few years many authors using Ly α systems have found high overdensities ($\delta > 50$) at high redshifts. Indeed, Steidel et al. (1998) using LBGs found a protocluster at $z \simeq 3.1$. Recently, Venemans et al. (2004), Miley et al. (2004), Ouchi et al. (2005) and Intema et al. (2006) found protoclusters of LAE galaxies at high redshifts ($z \simeq 4$). Also a cross correlation analysis has shown (Cooke et al. 2006) that the LBGs are associated with the DLA systems at the same redshift, a fact which implies that perhaps the LBGs trace the same parent system as the DLA systems (Cooke et al. 2005).

4 CONCLUSIONS

We have studied the clustering properties of the H I 21-cm emission line sources from HIPASS catalogue in redshift space. Modelling two-point correlation function as a power law, $\xi(r) = (r_0/r)^\gamma$, we find $r_0 = 3.3 \pm 0.3 h^{-1}$ Mpc and $\gamma = 1.38 \pm 0.24$. Fixing the slope to its universal value $\gamma = 1.8$, we obtain $r_0 = 3.2 \pm 0.2 h^{-1}$ Mpc.

Comparing the measured spatial correlation function for the H I selected galaxies with the theoretical predictions of the preferred Λ CDM cosmological model ($\Omega_m = 1 - \Omega_\Lambda = 0.26$) and a bias evolution model, we find that the present bias value is $b_0 \simeq 0.68$, suggesting that the H I selected galaxies are located in relatively underdense regions in the local universe, in agreement with previous studies (Grogan & Geller 1998). Using the derived value of the bias we determine the redshift-space distortion parameter, $K(\beta)$, and use it to derive the real-space correlation function, for which we have $r_0^e = 2.4 \pm 0.3 h^{-1}$ Mpc with $\gamma = 1.38$ and $r_0^e = 2.5 \pm 0.2 h^{-1}$ Mpc for $\gamma = 1.8$.

Dividing our H I galaxy sample into two richness subsamples we find: (i) for $M_{H I} \leq 2.475 \times 10^9 h^{-2} M_\odot$ that $r_0^e \simeq 1.8 h^{-1}$ Mpc (for $\gamma = 1.8$) and $b_0 \simeq 0.48$ while (ii) for $M_{H I} > 2.475 \times 10^9 h^{-2} M_\odot$ that $r_0^e \simeq 3.9 h^{-1}$ Mpc (for $\gamma = 1.8$) and $b_0 \simeq 0.94$. Therefore, the low-mass H I galaxies should trace low densities, due to very low biasing found while the high-mass H I galaxies should trace

average galaxy densities as optical galaxies do ($b_0 \simeq 1$). Indeed, we corroborate these results by correlating the H I galaxy positions with the IRAS–PSCz density field, smoothed over scales of $5 h^{-1}$ Mpc, to find that indeed the H I galaxies are typically found in negative overdensity regions ($\delta\rho/\rho_{PSCz} \lesssim 0$), even more so the low-mass H I galaxies. We conclude that such galaxies could be the typical population of galaxies in void-like regions.

Finally, we investigate the redshift evolution of the H I galaxy linear bias factor and we find that the antibias behaviour extends up to $z \leq 1.3$.

ACKNOWLEDGMENTS

We would like to thank the anonymous referee for his/her useful comments and suggestions. Dr Kovač now is a post-doctoral researcher at the Institut für Astronomie, ETH in Zürich.

In memory of Nikos Voglis. This paper is dedicated to Professor Nikos Voglis, who suddenly passed away at the age of 58, in his office at the Academy of Athens, while working on this manuscript. His colleagues and students will miss his exceptional intellect and enthusiasm.

REFERENCES

- Adelberger K. L., Steidel C. C., Gialalisko M., Dickinson M., Pettini M., Kellogg M., 1998, *ApJ*, 505, 18
- Arnouts S. et al., 2002, *MNRAS*, 329, 355
- Bagla J. S., 1998, *MNRAS*, 299, 424
- Bahcall N., Burgett W. S., 1986, *ApJ*, 300, L35
- Bardeen J. M., Bond J. R., Kaiser N., Szalay A. S., 1986, *ApJ*, 304, 15
- Barnes D. G. et al., 2001, *MNRAS*, 322, 486
- Basilakos S., Plionis M., 2001, *ApJ*, 550, 522
- Basilakos S., Plionis M., 2003, *ApJ*, 593, L61
- Basilakos S., Plionis M., 2005, *MNRAS*, 360, L35
- Basilakos S., Plionis M., 2006a, *ApJ*, 650, L1
- Basilakos S., Plionis M., 2006b, *MNRAS*, 373, 1112
- Benson A. J., Cole S., Frenk S. C., Baugh M. C., Lacey G. C., 2000, *MNRAS*, 311, 793
- Branchini E. et al., 1999, *MNRAS*, 308, 1
- Coil A. L., Newman J. A., Cooper M. C., Davis M., Faber S. M., Koo D. C., Willmer C. N. A., 2006, *ApJ*, 644, 671
- Colless M. et al., 2001, *MNRAS*, 328, 1039
- Cooke J. W., Wolfe A. M., Prochaska J. X., Gawiser E., 2005, *ApJ*, 621, 596
- Cooke J. W., Wolfe A. M., Gawiser E., Prochaska J. X., 2006, *ApJ*, 636, L9
- Davis M. et al., 2003, *Proc. SPIE*, 4834, 161, preprint (astro-ph/0209419)
- Doyle M. T. et al., 2005, *MNRAS*, 361, 34
- Efstathiou G., Bernstein G., Katz N., Tyson J. A., Guhathakurta P., 1991, *ApJ*, 380, L47
- Freedman W. L. et al., 2001, *ApJ*, 553, 47
- Fry J. N., 1996, *ApJ*, 461, 65
- Giovanelli R. et al., 2005a, *AJ*, 130, 2598
- Giovanelli R. et al., 2005b, *AJ*, 130, 2613
- Gonzalez-Solares E. A. et al., 2004, *MNRAS*, 352, 44
- Grogan N. A., Geller M. J., 1998, *ApJ*, 505, 506
- Guzzo L. et al., 2007, in Tresse L., Maurogordato S., Tran Thanh Van Y., eds, XXVth Astrophysics Moriond Meeting: ‘From Dark Halos to Light’, March 2006, *Proc. (Editions Frontieres)*, preprint (astro-ph/0701273)
- Hamilton A. J. S., 1992, *ApJ*, 385, L5
- Hawkins Ed. et al., 2003, *MNRAS*, 346, 78
- Intema H. T., Venemans B. P., Kurk J. D., Ouchi M., Kodama T. H., Rottgering J. A., Miley G. K., Overzier R. A., 2006, *A&A*, 456, 433
- Jing Y. P., Börner G., Suto Y., 2002, *ApJ*, 564, 15
- Kaiser N., 1984, *ApJ*, 284, L9
- Kashikawa N. et al., 2006, *ApJ*, 637, 631

⁵ For $z = 4.5$ we get $E(4.5) \simeq 6.64$ and $I(4.5) \simeq 6.41$.

- Kauffmann G., Golberg J. M., Diaferio A., White S. D. M., 1999, *MNRAS*, 307, 529
- Kirkman D., Tytler D., Suzuki N., O'Meara J. M., Lubin D., 2003, *ApJS*, 149, 1
- Lahav O. et al., 2002, *MNRAS*, 333, 961
- Lanzetta K. M., Wolfe A. M., Turnshek D. A., 1995, *ApJ*, 440, 435
- Le Fèvre O. et al., 2005, *A&A*, 439, 845
- Matarrese S., Coles P., Lucchin F., Moscardini L., 1997, *MNRAS*, 286, 115
- Meyer M. J. et al., 2004, *MNRAS*, 350, 1195
- Meyer M. J., Zwaan M. A., Webster R. L., Brown M. J. I., Staveley-Smith L., 2007, *ApJ*, 654, 702
- Miley G. K. et al., 2004, *Nat*, 427, 47
- Mo H. J., White S. D. M., 1996, *MNRAS*, 282, 347
- Mo H. J., White S. D. M., 2002, *MNRAS*, 336, 112
- Mo H. J., Jing Y. P., Börner G., 1992, *ApJ*, 392, 452
- Norberg P. et al., 2001, *MNRAS*, 328, 64
- Nusser M., Davis M., 1994, *ApJ*, 421, L1
- Olive K. A., Steigman G., Walker T. P., 2000, *Phys. Rep.*, 333, 389
- Oliver S. et al., 2000, *MNRAS*, 316, 749
- Ouchi M. et al., 2005, *ApJ*, 620, L1
- Peacock A. J., Dodds S. J., 1994, *MNRAS*, 267, 1020
- Peebles P. J. E., 1993, *Principles of Physical Cosmology*. Princeton Univ. Press, Princeton, NJ
- Peebles P. J. E., Ratra B., 2003, *Rev. Mod. Phys.*, 75, 559
- Peroux C., Storrie-Lombardi L. J., McMahon R. G., Irwin M., Hook I. M., 2001, *AJ*, 121, 1799
- Plionis M., Basilakos S., 2001, *MNRAS*, 327, L32
- Prochaska J. X., 1999, *ApJ*, 511, L71
- Riess A. G. et al., 2004, *ApJ*, 607, 665
- Ryan-Weber E. V., 2006, *MNRAS*, 367, 1251
- Saunders W. et al., 2000, *MNRAS*, 317, 55
- Spergel D. N. et al., 2003, *ApJS*, 148, 175
- Spergel D. N. et al., 2006, *ApJ*, in press (astro-ph/0603449)
- Steidel C. C., Adelberger L. K., Dickinson M., Giavalisco M., Pettini M., Kellogg M., 1998, *ApJ*, 492, 428
- Sugiyama N., 1995, *ApJS*, 100, 281
- Tegmark M., Peebles P. J. E., 1998, *ApJ*, 500, L79
- Tegmark M. et al., 2004, *Phys. Rev. D.*, 69, 3501
- Tonry J. L. et al., 2003, *ApJ*, 594, 1
- Vasilyev N. L., Baryshev Yu. V., Sylos Labini F., 2006, *A&A*, 447, 431
- Venemans B. P. et al., 2004, *A&A*, 427, L17
- Verde L., 2002, *MNRAS*, 335, 432
- Viegas S. M., 1995, *MNRAS*, 276, 268
- Vladilo G., Centurion M., Bonifacio P., Howk J. C., 2001, *ApJ*, 557, 1007
- Wolfe A. M., Turnshek D. A., Smith H. E., Cohen R. D., 1986, *ApJS*, 61, 249
- Wolfe A. M., Gawiser E., Prochaska J. X., 2005, *ARA&A*, 43, 861
- Wong O. I. et al., 2006, *MNRAS*, 371, 1855
- York D. G., 2000, *AJ*, 120, 1579
- Zehavi I. et al., 2002, *ApJ*, 571, 172
- Zwaan M. A. et al., 2004, *MNRAS*, 350, 1210
- Zwaan M. A., Meyer M. J., Staveley-Smith L., Webster R. L., 2005a, *MNRAS*, 359, L30
- Zwaan M. A., van der Hulst J. M., Briggs H. F., Verheijen M. A. W., Ryan-Weber E. V., 2005b, 364, 1467
- Zwaan M. A., Prochaska J. X., 2006, *ApJ*, 643, 675

This paper has been typeset from a \LaTeX file prepared by the author.